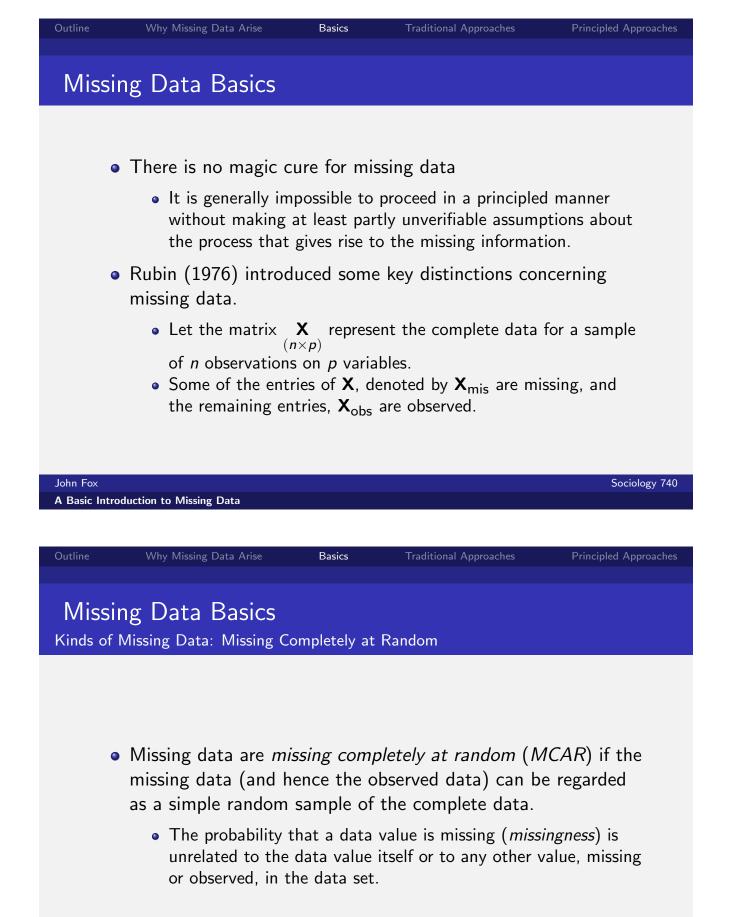
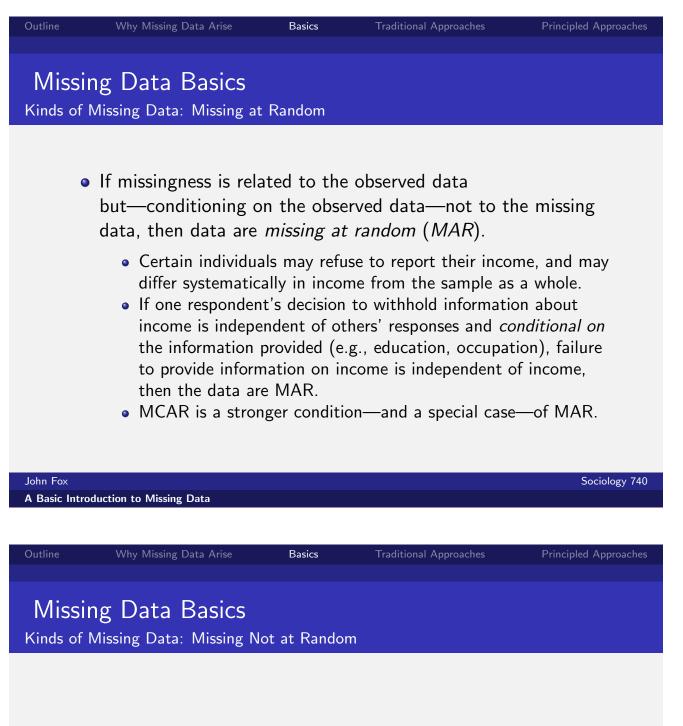
Outline	Why Missing Data Arise	Basics	Traditional Approaches	Principled Approaches				
	A Basic Intro	oductio	n to Missing D	Data				
		John	Fox					
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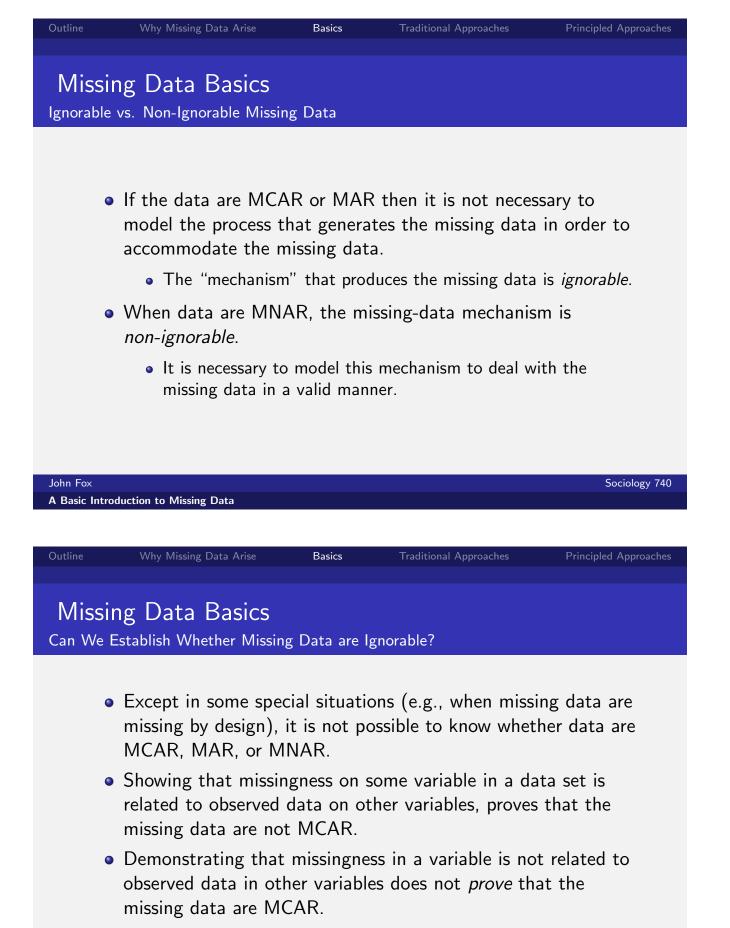
	Traditional Approaches	Principled Approaches					
se							
esponse.							
 In a survey, certain respondents may be unreachable or may refuse to participate. 							
		pecific					
tion or pro	ocessing.						
nay fail to a	sk a question of a surv	/ey					
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Basics	Traditional Approaches	Principled Approaches					
se							
e built into	o the design of a stu	dy.					
 Particular questions may be asked only of a random subset of respondents. 							
 Some data values in a study may be censored. 							
a study n	nay be <i>censored</i> .						
t-history) a	nay be <i>censored</i> . nalysis, the focal event ome subjects before the	• -					
	ain responde bate. ts may not y refuse to tion or pro hay fail to a Basics SE SE	ain respondents may be unreachable bate. ts may not know the answers to say refuse to respond to them. tion or processing. hay fail to ask a question of a survey Basics Traditional Approaches					

• A survey respondent who has no children cannot report their ages.

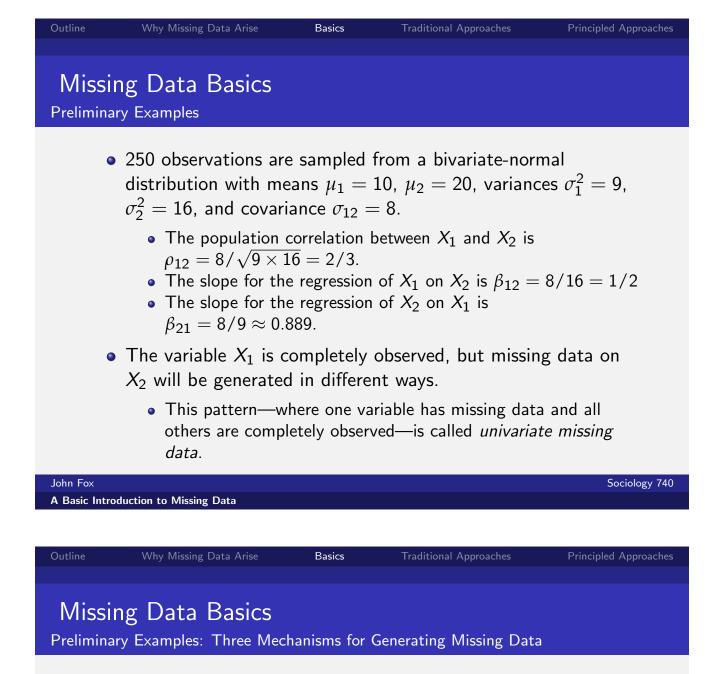




- If missingness is related to the missing values themselves even when the information in the observed data is taken into account, then missing data are *missing not at random* (*MNAR*).
 - If conditional on all of the observed data, individuals with higher incomes are more likely than others to withhold information about their incomes, then the missing income data are MNAR.



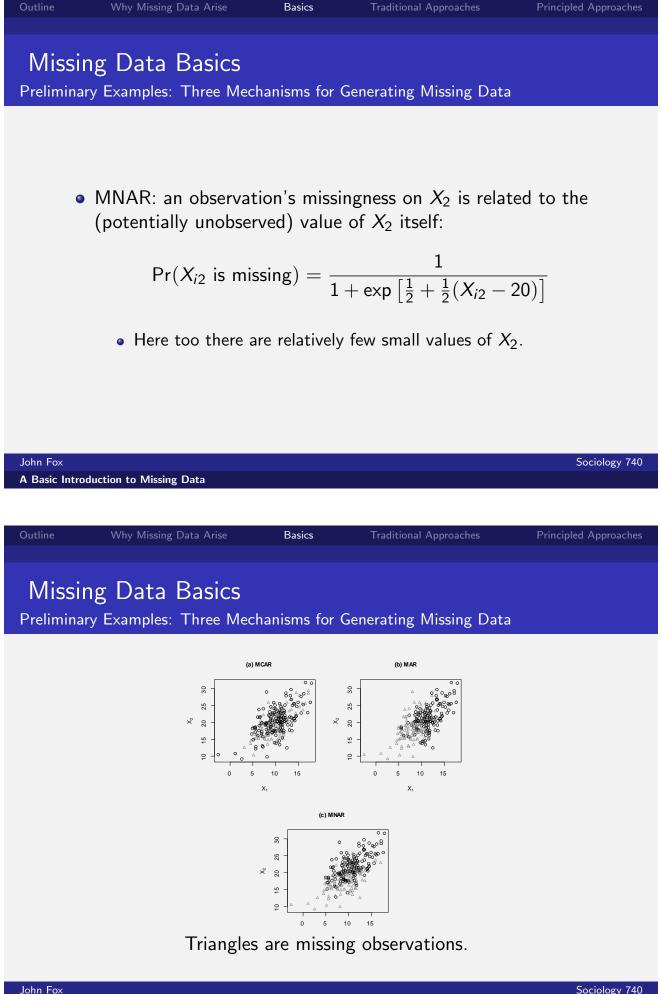
• Non-respondents may be differentiated from respondents in some *unobserved* manner.



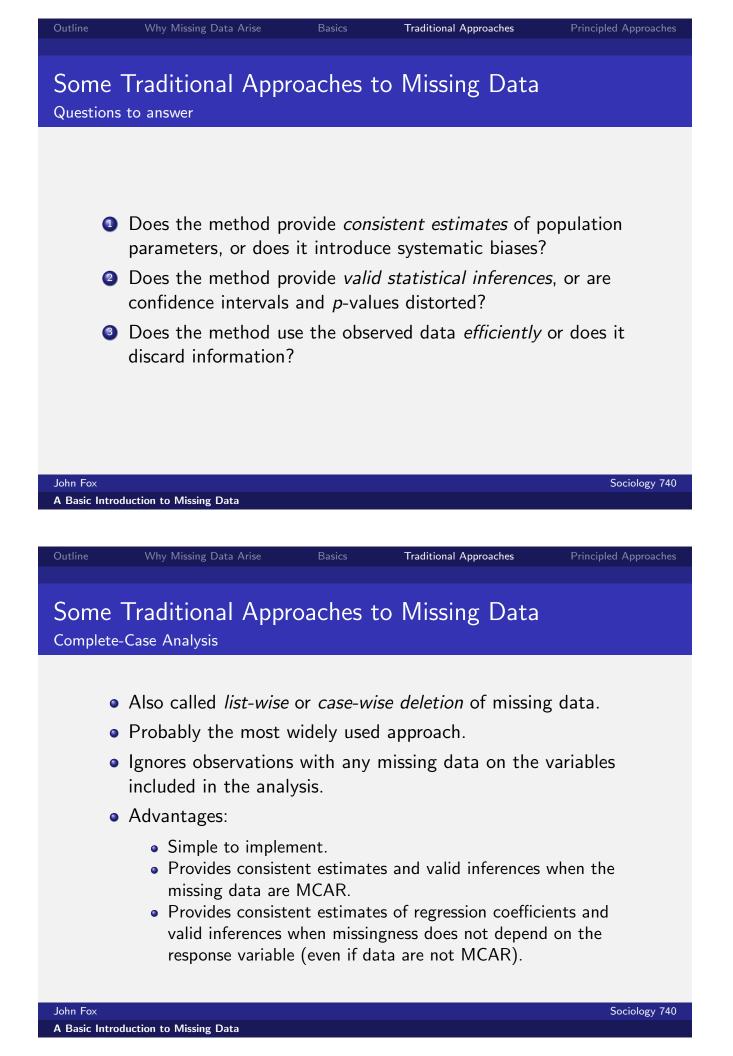
- MCAR: 100 of the observations on X₂ are selected at random and set to missing.
- MAR: an observation's missingness on X₂ is related to its (observed) value of X₁:

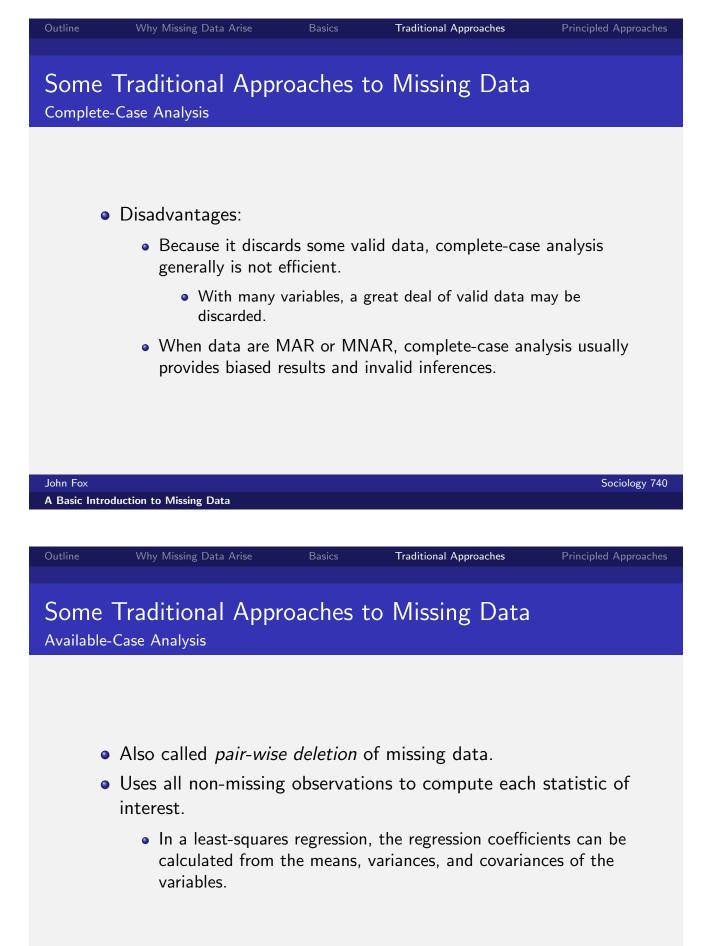
$$\Pr(X_{i2} \text{ is missing}) = \frac{1}{1 + \exp\left[\frac{1}{2} + \frac{2}{3}(X_{i1} - 10)\right]}$$

- The logistic regression coefficients were calibrated so that approximately 100 observations will have missing data on X₂, with the probability that X₂ is missing declining as X₁ grows.
- Because X₁ and X₂ are positively correlated, there are relatively few small values of X₂.



A Basic Introduction to Missing Data





Some Traditional Approaches to Missing Data

Available-Case Analysis

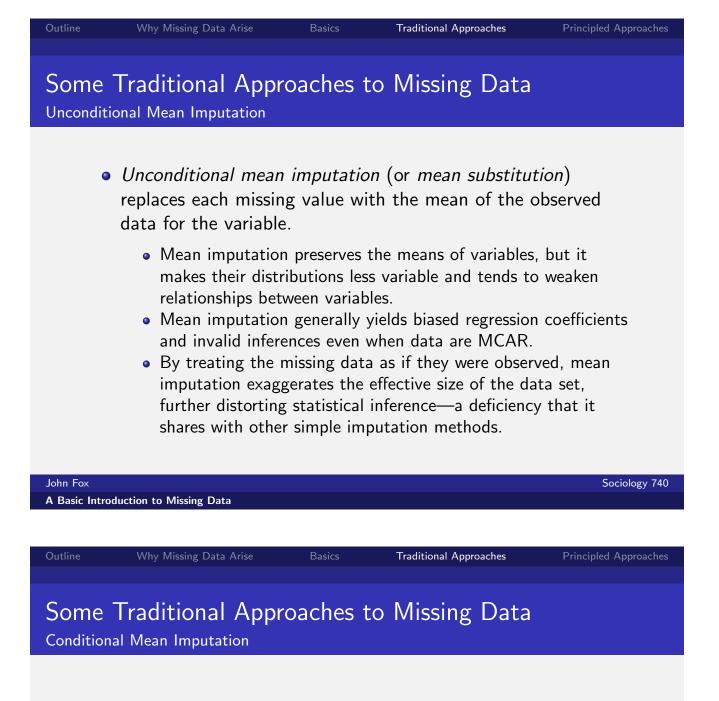
• Problems:

• Available case analysis appears to use more information than complete-case analysis, but estimators based on available cases can be *less* efficient than those based on complete cases.

Traditional Approaches

- By basing different statistics on different subsets of the data, available-case analysis can lead to nonsensical results, such as correlations outside the range from -1 to +1.
- Except in simple cases, such as linear least-squares regression, it is not obvious how to apply the available-case approach.
- Available-case analysis generally provides biased estimates and invalid inferences when data are MAR or MNAR.

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Imputatio	on			
	Replacing missing v	alues with r	plausible <i>imputed</i> va	lues
•				1005.
	Ŭ	•	a set is then analyzed	using
	standard method	ls.		



- Conditional-mean imputation replaces missing data with predicted values, obtained, for example, from a regression equation (*regression imputation*).
 - The imputed observations tend to be less variable than real data, because they lack residual variation.
 - Another problem is that we have failed to account for uncertainty in the estimation of the regression coefficients used to obtain the imputed values.
 - Regression imputation improves on unconditional mean imputation, but it generally provides biased estimates and invalid inferences even for missing data that are MCAR.



Parameter	Complete	Mean	Regr.	Multiple
Farameter	Cases	Imput.	Imput.	Imput.
	Mean	Estimate (R	MSE)	
$\mu_1 = 10$	11.476	10.001	10.001	10.001
	(1.489)	(0.189)	(0.189)	(0.189)
$\mu_2 = 20$	21.222	21.322	20.008	20.008
	(1.355)	(1.355)	(0.326)	(0.344)
$\beta_{12} = 0.5$	0.391	0.391	0.645	0.498
	(0.117)	(0.117)	(0.151)	(0.041)
$\beta_{21} = 0.889$	0.891	0.353	0.891	0.890
	(0.100)	(0.538)	(0.100)	(0.106)

Some Traditional Approaches to Missing Data

Application to the Illustrative MAR Data: Based on A Simulation With 1000 Samples

Parameter	Complete	Mean	Regr.	Multiple
Farameter	Cases	Imput.	Imput.	Imput.
	ConfInterval Coverage		(Mean Inte	rval Width)
μ_1	0	.951	.951	.951
	(0.792)	(0.750)	(0.750)	(0.746)
μ_2	.005	0	.823	.947
	(1.194)	(0.711)	(0.881)	(1.451)
β_{12}	.304	.629	.037	.955
	(0.174)	(0.246)	(0.140)	(0.175)
β_{21}	.953	0	.661	.939
	(0.396)	(0.220)	(0.191)	(0.463)

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Principled Approaches to Missing Data

Maximum-Likelihood Estimation

- The method of maximum likelihood can be applied to parameter estimation in the presence of missing data.
 - Doing so requires making assumptions about the distribution of the complete data and about the process producing missing data.
 - If the assumptions hold, then the resulting maximum-likelihood estimates have their usual optimal properties, such as consistency and asymptotic efficiency.
- Let p(X; θ) = p(X_{obs}, X_{mis}; θ) represent the joint probability-density for the complete data X.
 - Rubin (1976) showed that the maximum-likelihood estimate $\hat{\theta}$ of θ can be obtained from the marginal distribution of the observed data, *if data are MAR*.

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- Bayesian multiple imputation (MI) is a flexible and general method for dealing with data that are missing at random.
- Like maximum-likelihood estimation, MI begins with a specification of the distribution of the complete data (assumed to be known except for a set of parameters to be estimated from the data).

Bayesian Multiple Imputation

- The essential idea of multiple imputation is to reflect the uncertainty associated with missing data by imputing *several* values for each missing value, each imputed value drawn from the *predictive distribution* of the missing data, and therefore producing not one but several completed data sets.
 - Standard methods of statistical analysis are then applied in parallel to the completed data sets.
 - Parameters of interest are estimated along with their standard errors for each imputed data set.
 - Estimated parameters are averaged across completed data sets.
 - Standard errors are combined across imputed data sets, taking into account the variation among the estimates in the several data sets, thereby capturing the added uncertainty due to having to impute the missing data.

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- A multivariate-normal model for the complete data is both relatively simple and useful in applications.
 - Because the model assumed to describe the complete data is used just to obtain imputed values for the missing data, results produced by MI are usually not sensitive to the assumption of multivariate normality.
 - But there are some pitfalls to be avoided (discussed below).
- The details of methods for drawing multiple imputations from the multivariate-normal model are beyond the scope of this presentation.

Bayesian Multiple Imputation: Rubin's Rules

- Having obtained g completed data sets, we produce g sets of regression coefficients, $B_0^{(I)}$, $B_1^{(I)}$, ..., $B_k^{(I)}$, and coefficient standard errors, $SE(B_0^{(I)})$, $SE(B_1^{(I)})$, ..., $SE(B_k^{(I)})$, for $l = 1, \ldots, g$.
- Rubin (1987) provides simple rules for combining information across multiple imputations, valid as long as the sample size is sufficiently large for the separate estimates to be approximately normally distributed.
 - Point estimates of the population regression coefficients are obtained by averaging across imputations:

$$\widetilde{\beta}_j \equiv \frac{\sum_{l=1}^{g} B_j^{(l)}}{g}$$

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Bayesian Multiple Imputation: Rubin's Rules

 Standard errors of the estimated coefficients are obtained by combining information about within and between-imputation variation in the coefficients:

$$\widetilde{\mathsf{SE}}(\widetilde{\beta}_j) \equiv \sqrt{V_j^{(W)} + \frac{g+1}{g}V_j^{(B)}}$$

where the within-imputation component is

$$V_{j}^{(W)} \equiv \frac{\sum_{l=1}^{g} SE^{2}(B_{j}^{(l)})}{g} \text{ and the between-imputation component}$$
is $V_{j}^{(B)} \equiv \frac{\sum_{l=1}^{g} (B_{j}^{(l)} - \tilde{\beta}_{j})^{2}}{g-1}.$

Bayesian Multiple Imputation: Rubin's Rules

• Inference based on $\widetilde{\beta}_j$ and $\widetilde{SE}(\widetilde{\beta}_j)$ uses the *t*-distribution, with degrees of freedom

$$extsf{df}_{j} = (g-1) \left(1 + rac{g}{g+1} imes rac{V_{j}^{(W)}}{V_{j}^{(B)}}
ight)^{2}$$

• For example, to construct a 95-percent confidence interval for β_j ,

$$\beta_j = \widetilde{\beta}_j \pm t_{.025, df_j} \widetilde{\mathsf{SE}}(\widetilde{\beta}_j)$$

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$$\widehat{\gamma}_j = \frac{R_j}{R_j + 1}$$

where

$$R_j \equiv rac{g+1}{g} imes rac{V_j^{(B)}}{V_j^{(W)}}$$

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efficient maximum-likelihood estimator is

$$\mathsf{RE}(\widetilde{\beta}_j) = rac{g}{g + \gamma_j}$$

- If the number of imputations g is very large, MI is as efficient as ML.
- Even when the rate of missing information is high and the number of imputations modest, the relative efficiency of the MI estimator hardly suffers.
 - When $\gamma_i = 0.5$ and g = 5, then

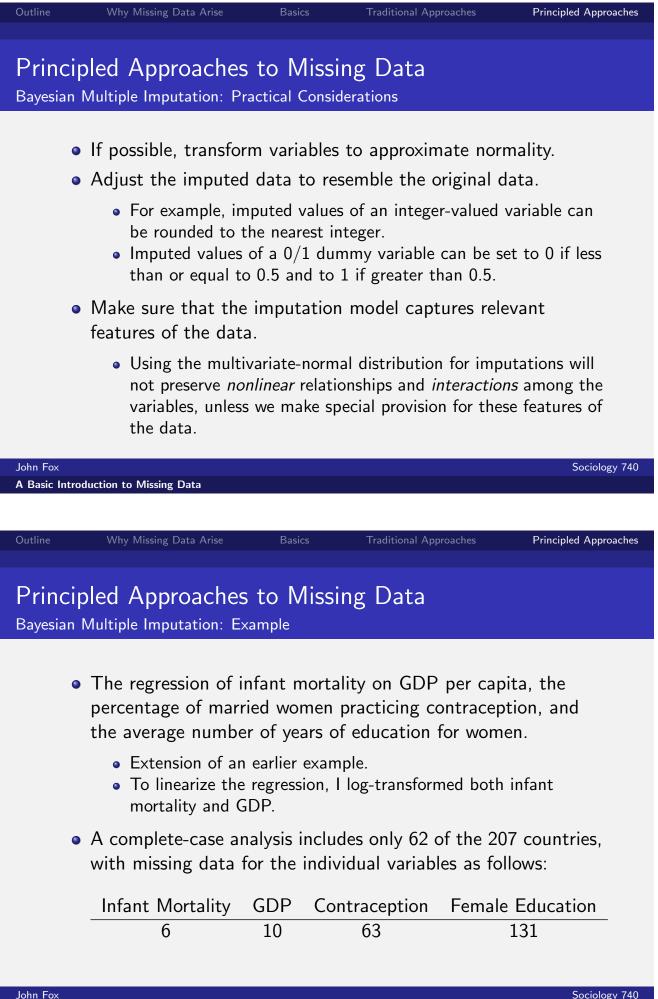
$$\mathsf{RE}(\widetilde{eta}_j) = 5/(5+0.5) = 0.91$$
, and $\sqrt{\mathsf{RE}(\widetilde{eta}_j)} = 0.95$.

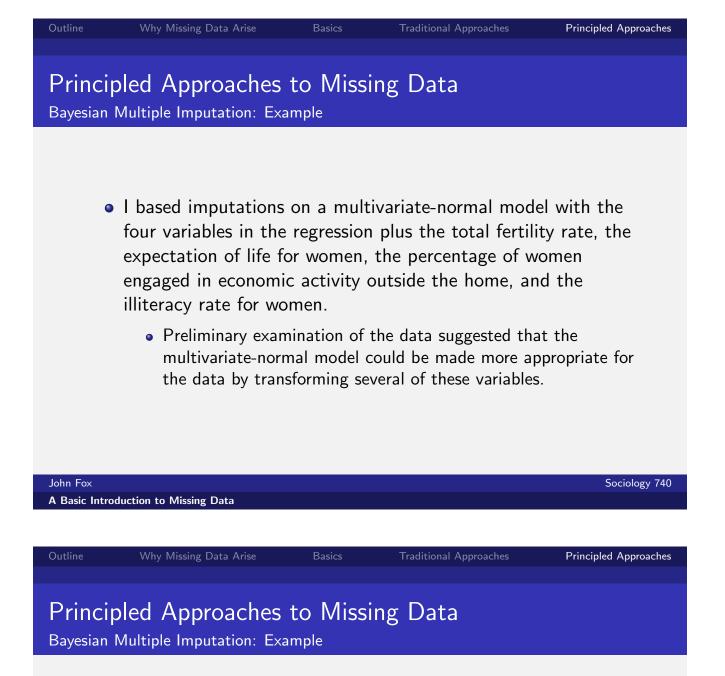
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Principled Approaches to Missing Data

Bayesian Multiple Imputation: Practical Considerations

- Multiple imputation cannot preserve features of the data that are not represented in the imputation model.
 - It is important to insure that the imputation model is consistent with the intended analysis.
- Try to include variables in the imputation model that make the assumption of ignorable missingness reasonable.
 - Think of imputation as a pure prediction problem.
 - Finding variables that are highly correlated with a variable that has missing data will likely improve the quality of imputations, as will variables that are related to missingness.
 - Use all relevant variables, even ones not used in the substantive analysis (an "inclusive" strategy).
 - There is nothing wrong in using the response variable to help impute missing values of explanatory variables.





 Means and standard deviations of the variables are as follows, based on complete cases and on ML assuming ignorable missing data:

	log _e Inf.Mort.	$\log_e \text{GDP}$	DP Contra. Fema	
	Estimates based on Complete Case			ses
Mean	3.041	8.151	50.90	11.30
SD	(1.051)	(1.703)	(23.17)	(3.55)
Maximum-Likelihood Estimates				
Mean	3.300	7.586	44.36	10.16
SD	(1.022)	(1.682)	(24.01)	(3.51)

Bayesian Multiple Imputation: Example

- Using Schafer's (1997) data-augmentation method, and employing the multivariate-normal model, I obtained imputations for 10 completed data sets.
- Results:

	Intercept	$\log_e \text{GDP}$	Contra.	Female Ed.
		Complete	-Case Analysi	is
Bj	6.88	-0.294	-0.0113	-0.0770
$SE(B_j)$	(0.29)	(0.058)	(0.0042)	(0.0338)
	I	Multiple-Im	putation Anal	lysis
\widetilde{eta}_j	6.57	-0.234	-0.00953	-0.105
$\begin{vmatrix} \beta_j \\ \widetilde{SE}(\widetilde{\beta}_j) \end{vmatrix}$	(0.18)	(0.049)	(0.00294)	(0.033)
Miss. Inf. $\widehat{\gamma}_j$	0.20	0.61	0.41	0.69

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